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Automated Insulin Delivery Systems in Children and Adolescents With Type 1 Diabetes: A Systematic Review and Meta-analysis of Outpatient Randomized Controlled Trials

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Automated Insulin Delivery Systems in Children and Adolescents With Type 1 Diabetes: A Systematic Review and Meta-analysis of Outpatient Randomized Controlled Trials

Design and Methods

- PubMed, Embase, the Cochrane Library, and ClinicalTrials.gov were searched until 4 May 2023.
- We included randomized controlled trials that compared automated insulin delivery (AID) systems with conventional insulin therapy in outpatient children and adolescents with type 1 diabetes (T1D)
- Percent time in range (TIR, 3.9–10 mmol/L), time below range (TBR, <3.9 mmol/L), and time above range (TAR, >10 mmol/L) were extracted. Data were summarized as mean differences with 95% CI.

Results

- 25 trials (1,345 participants) were included in the meta-analysis.
- AID systems were associated with an increased percentage of TIR (high certainty).
- The favorable effect was consistent whether AID was used over 3 months or 6 months.
- AID systems had a favorable effect on the proportion of TBR (low certainty) or TAR (high certainty), compared with control treatment.

| Outcomes | | | | | Mean |
|---------------------|--------|---|---|---|-------------------------|
| and subgroups | Trials | | | | difference (95% CI) |
| 3.9–10 mmol/L | | | | | |
| All comparisons | 25 | | | | 11.38 (9.01, 13.76) |
| Unsupervised | 18 | | | + | 9.88 (8.05, 11.72) |
| Supervised | 7 | | | | - 15.25 (5.55, 24.96) |
| Duration ≥ 3 months | 9 | | | + | 10.46 (8.71, 12.20) |
| Duration ≥ 6 months | 5 | | | | 10.87 (7.11, 14.63) |
| <3.9 mmol/L | | | | | |
| All comparisons | 23 | | | | -0.59 (-1.02, -0.15) |
| Unsupervised | 17 | | + | | -0.23 (-0.67, 0.21) |
| Supervised | 6 | | • | | -2.03 (-3.09, -0.96) |
| >10 mmol/L | | | | | |
| All comparisons | 18 | | | | -12.19 (-14.65, -9.73) |
| Unsupervised | 14 | - | | | -9.74 (-11.60, -7.88) |
| Supervised | 4 | | | | -19.32 (-26.64, -12.00) |

AID systems are more effective than conventional insulin therapy for children and adolescents with T1D in outpatient settings. The favorable effect is consistent both in the short term and long term.

ARTICLE HIGHLIGHTS

- The findings presented in this study are based on data from 25 randomized controlled trials of automated insulin delivery (AID) systems involving 1,345 children and adolescents with type 1 diabetes.
- High-certainty evidence indicates that the use of AID systems results in an 11.38% increased time in target blood glucose range, equivalent to 164 min/day.
- This favorable effect was observed consistently when the AID was used continuously over 3 and 6 months.



Automated Insulin Delivery Systems in Children and Adolescents With Type 1 Diabetes: A Systematic Review and Meta-analysis of Outpatient Randomized Controlled Trials



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BACKGROUND

The glycemic control of automated insulin delivery (AID) systems in outpatient children and adolescents with type 1 diabetes (T1D) has not been systematically evaluated.

PURPOSE

META-ANALYSIS

To evaluate the efficacy and safety of AID systems in children and adolescents in outpatient settings.

DATA SOURCES

PubMed, Embase, the Cochrane Central Register of Controlled Trials, and ClinicalTrials.gov were searched until 4 May 2023. This study was registered with PROSPERO (2023, CRD42023395252).

STUDY SELECTION

Randomized controlled trials that compared AID systems with conventional insulin therapy in outpatient children and adolescents with T1D and reported continuous glucose monitoring outcomes were selected.

DATA EXTRACTION

Percent time in range (TIR) (3.9–10 mmol/L), time below range (TBR) (<3.9 mmol/L), and time above range (TAR) (>10 mmol/L) were extracted. Data were summarized as mean differences (MDs) with 95% CIs.

DATA SYNTHESIS

Twenty-five trials (1,345 participants) were included in the meta-analysis. AID systems were associated with an increased percentage of TIR (MD, 11.38% [95% CI 9.01–13.76], P < 0.001; high certainty). The favorable effect was consistent whether AID was used over 3 months (10.46% [8.71–12.20]) or 6 months (10.87% [7.11–14.63]). AID systems had a favorable effect on the proportion of TBR (-0.59% [-1.02 to -0.15], P = 0.008; low certainty) or TAR (-12.19% [-14.65 to -9.73], P < 0.001; high certainty) compared with control treatment.

LIMITATIONS

Substantial heterogeneity was observed in most analyses.

CONCLUSIONS

AID systems are more effective than conventional insulin therapy for children and adolescents with T1D in outpatient settings. The favorable effect is consistent both in the short term and long term. ¹Central Laboratory, Peking University Binhai Hospital, Tianjin, China

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Type 1 diabetes (T1D) is a chronic autoimmune disease characterized by insulin deficiency and resultant hyperglycemia (1). According to the International Diabetes Federation Diabetes Atlas, an estimated 1.21 million children and adolescents aged <20 years have T1D, with 149,500 new cases diagnosed annually worldwide in 2021 (2). The American Diabetes Association now recommends that glycated hemoglobin (HbA_{1c}) level be kept at ${<}7\%$ (3), which corresponds to a time in range (TIR) >65-75% as measured by continuous glucose monitoring (4,5). Despite increased use of insulin pump therapy and continuous glucose monitoring, <10% of children and adolescents aged \leq 17 years in the U.S. reach the international target (6). T1D in children often leads to a high management burden and reduced quality of life for the whole family (7).

Automated insulin delivery (AID), a novel technology, is a hybrid closed-loop system that automatically administers an insulin dose based on glucose sensor readings and a dosing algorithm (8,9). Three previous meta-analyses have concluded that the AID systems could improve glucose control compared with conventional insulin pump therapy in outpatients with T1D, but most of the included trials recruited adults or mixed populations (10-12). Another metaanalysis showed that AID systems were superior to the standard sensor-augmented pump treatment of T1D in children and adolescents, but most of the trials included were in inpatient settings, and some were not randomized (13). Finally, a recent meta-analysis summarized evidence from published trials AID systems in children and adolescents with T1D (14). In recent years, some important randomized controlled trials (RCTs) assessing AID systems in young people were also published. Therefore, we performed a meta-analysis of AID systems for glycemic control in children and adolescents with T1D compared with conventional insulin therapy in outpatient settings.

RESEARCH DESIGN AND METHODS

Data Sources and Searches

We conducted this meta-analysis according to the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) guidelines (15). The protocol was registered with the International prospective register of systematic reviews (PROSPERO) (2023, CRD42023395252 [https://www.crd.york.ac.uk/prospero/ display_record.php?RecordID=395252]). We searched for literature published in PubMed, Embase, and the Cochrane Central Register of Controlled Trials and gray literature from ClinicalTrials.gov until 4 May 2023. Keywords included "type 1 diabetes," "artificial pancreas," "closed loop," "children," and "adolescent." The detailed search strategy is provided in the Supplementary Material. Our search was restricted to articles published in English. Additionally, we identified references by searching the reference lists of included studies and relevant reviews.

Study Selection

We included RCTs in children and adolescents aged \leq 25 years with T1D seen in the outpatient setting, irrespective of trial design (parallel or crossover) or timing of intervention (24 h or overnight), which compared AID systems with conventional insulin therapy. The outpatient setting was defined as the participant's home, hotel. diabetes camp, or research house. The conventional insulin therapy included multiple daily insulin injection (MDI), continuous subcutaneous insulin infusion (CSII), and sensor-augmented pump (SAP) with or without low glucose suspend. We excluded studies with very short intervention periods (<3 days). Of note, only studies reporting continuous glucose monitoring outcomes were included.

The primary outcome was the percentage of TIR, i.e., percentage of the total duration of the intervention that blood glucose was within target range (3.9–10 mmol/L). Secondary outcomes included time below range (TBR) (<3.9 mmol/L), TBR (<3.0 mmol/L), time above range (TAR) (>10 mmol/L), TAR (>13.9 mmol/L), severe hypoglycemic event (as defined in each individual study), and diabetic ketoacidosis.

Data Extraction and Quality Assessment

Two authors reviewed the titles and abstracts independently to identify eligible studies that met prespecified inclusion criteria and extracted the data. When consensus was lacking, a third reviewer was consulted. Study characteristics (e.g., year of publication, study design, sample size), intervention and comparator characteristics, patient characteristics (e.g., age, sex, diabetes duration, baseline HbA_{1c}), and outcomes were extracted. The risk of bias of RCTs was assessed using the Cochrane Collaboration's tool (16). If the study reported both 24-h and overnight results, we extracted both results. The quality of evidence for each outcome was evaluated using the Grading of Recommendations Assessment, Development and Evaluation (GRADE) framework (17).

Data Synthesis and Analysis

We conducted DerSimonian and Laird random-effects meta-analysis when data were available for at least two studies (18). TIR. TBR. and TAR were analyzed as mean differences (MDs) with 95% Cls. We only performed narrative descriptive synthesis for severe hypoglycemic events and diabetic ketoacidosis. Medians were assumed to equal means, and SD was calculated as the interguartile range/1.35, which was recommended by the Cochrane Collaboration (16). We combined data from both parallel and crossover trials. If crossover trials did not report the mean and SE of the paired differences, we planned a priori to analyze all studies using group means and SDs, assuming a correlation coefficient of 0.5. For studies comparing both dual-hormone and singlehormone AIDs with conventional insulin therapy in a three-way crossover design, we included two comparisons (dualhormone vs. control and single-hormone vs. control) in the meta-analysis. Statistical heterogeneity among the studies was assessed using χ^2 test and I^2 statistics. I^2 values of 25%, 50%, and 75% have been suggested to be indicators of low, moderate, and high heterogeneity, respectively (19). Publication bias was assessed using a funnel plot for the primary outcome and Egger test (20,21), where P < 0.05 indicates the presence of publication bias.

For the primary outcome and two secondary outcomes (TBR [<3.9 mmol/L] and TAR [>10 mmol/L]), we conducted a prespecified subgroup analysis based on timing of the intervention (24 h vs. overnight) and mean age (<14 vs. \geq 14 years) and conducted a post hoc subgroup analysis based on setting (unsupervised home vs. supervised diabetes camp and hotel), AID system (single vs. dual hormone), study design (parallel vs. crossover), study duration (<1 month vs. >1 month), and comparator (SAP vs. CSII/MDI). The P value for the difference was calculated using random-effects meta-regression, and a difference between the estimates of these

| Study | Patients | Mean difference ts Timing of intervention Study duration with 95% Cl | | nce I | Weight (%) | | |
|--|----------------------------|---|-----------|----------|----------------|--------|------|
| Unsupervised | | | | | | | |
| Sharifi 2016 | 12 | Overnight | 4 days | | -0.90 [-5.95, | 4.15] | 3.95 |
| Russell 2018 | 20 | 24 h | 5 days | | 0.00 [-4.38, | 4.38] | 4.14 |
| Ware 2022b | 133 | 24 h | 6 months | | 6.70 [2.15, | 11.25] | 4.09 |
| Abraham 2021 | 135 | 24 h | 6 months | | 6.70 [2.65, | 10.75] | 4.22 |
| Kariyawasam 2022 | 21 | 24 h | 6 weeks | - | 7.51 [4.58, | 10.44] | 4.49 |
| Ware 2022a | 74 | 24 h | 16 weeks | | 8.70 [7.45, | 9.95] | 4.75 |
| Thabit 2015 | 25 | Overnight | 12 weeks | - | 8.90 [5.95, | 11.85] | 4.48 |
| Forlenza 2018 | 28 | Overnight | 21 days | | 10.00 [2.56, | 17.44] | 3.27 |
| Messer 2022 | 165 | 24 h | 13 weeks | - | 10.00 [7.00, | 13.00] | 4.47 |
| von dem Berge 2022 | 38 | 24 h | 8 weeks | | 10.50 [6.67, | 14.33] | 4.28 |
| Breton 2020 | 101 | 24 h | 16 weeks | | 11.00 [7.50, | 14.50] | 4.36 |
| Hovorka 2014 | 16 | Overnight | 3 weeks | | 12.00 [6.92, | 17.08] | 3.94 |
| Wadwa 2023 | 102 | 24 h | 13 weeks | - | 12.40 [9.50, | 15.30] | 4.49 |
| lsganaitis 2021 | 63 | 24 h | 6 months | | 13.00 [9.50, | 16.50] | 4.36 |
| Boughton 2022 | 97 | 24 h | 24 months | | 14.00 [6.50, | 21.50] | 3.25 |
| McVean 2023 | 113 | 24 h | 12 months | | 16.00 [10.00, | 22.00] | 3.68 |
| Tauschmann 2016b | 12 | 24 h | 3 weeks | | 18.80 [13.26, | 24.34] | 3.81 |
| Tauschmann 2016a | 12 | 24 h | 7 days | | 19.00 [12.26, | 25.74] | 3.47 |
| Heterogeneity: $\tau^2 = 10.8$ | $1, I^2 = 77.339$ | %, $H^2 = 4.41$ | | • | 9.88 [8.05, | 11.72] | |
| Test of $\theta_i = \theta_j$: Q(17) = 74 | 4.99, p = 0.0 | 0 | | | | | |
| Supervised | | | | | | | |
| Del Favero 2016 | 30 | 24 h | 3 days | | -6.30 [-10.75, | -1.85] | 4.12 |
| Breton 2017 | 32 | 24 h | 5 days | | 6.60 [-4.21, | 17.41] | 2.41 |
| Russell 2014 | 32 | 24 h | 5 days | | 11.40 [7.16, | 15.64] | 4.17 |
| Ly 2014 | 20 | Overnight | 6 days | | 12.00 [-1.44, | 25.44] | 1.89 |
| Haidar 2015a | 33 | Overnight | 3 days | | 16.00 [8.42, | 23.58] | 3.23 |
| Russell 2016 | 19 | 24 h | 5 days | | 23.00 [17.55, | 28.45] | 3.84 |
| DeBoer 2017 | 12 | 24 h | 3 days | | 26.20 [20.91, | 31.49] | 3.88 |
| Haidar 2015b | 33 | Overnight | 3 days | | 33.00 [24.41, | 41.59] | 2.95 |
| Heterogeneity: $\tau^2 = 180.$ | $03, I^2 = 94.71$ | 1% , $H^2 = 18.91$ | | | 15.25 [5.55, | 24.96] | |
| Test of $\theta_i = \theta_j$: Q(7) = 132 | 2.35, p = 0.0 | 0 | | | | | |
| Overall | | | | • | 11.38 [9.01, | 13.76] | |
| Heterogeneity: $\tau^2 = 30.4$ | 4, I ² = 88.519 | %, $H^2 = 8.70$ | | | | | |
| Test of $\theta_i = \theta_j$: Q(25) = 2 | 17.54, p = 0. | 00 | | | | | |
| | | | F | | | | |

Figure 1—Forest plot for TIR (3.9–10 mmol/L) by study included in the meta-analysis. Full reference citations are available in the Supplementary Material.

subgroups was considered significant if $P_{\text{interaction}} < 0.10$ (22). All the analysis were performed with Stata 17 statistical software.

Data and Resource Availability

The data sets used and/or analyzed during the current study are available from the corresponding author upon reasonable request.

RESULTS

Characteristics of Included Studies

The systematic literature search initially identified 1,038 records. After excluding duplicates and irrelevant articles, 90

articles were evaluated in full text for eligibility (Supplementary Fig. 1). Finally, 25 RCTs (1,345 participants) were included in the present meta-analysis (full reference citations are available in the Supplementary Material). Ten trials were parallel, and the remaining 15 trials were crossover. All the studies were published

| Table 1—Subgroup analyse | s for primary and seco | ndary outcomes | | | |
|--------------------------------------|------------------------|---|---------|--------------------|--------------------------|
| Subgroup | Comparisons, n | RD (95% CI) | Р | l ² , % | P _{interaction} |
| TIR (3.9–10.0 mmol/L) | | | | | |
| All comparisons | 26 | 11.38 (9.01–13.76) | <0.001 | 89 | |
| Setting | | | | | |
| Unsupervised | 18 | 9.88 (8.05–11.72) | <0.001 | 77 | 0.156 |
| Supervised | 8 | 15.25 (5.55–24.96) | 0.002 | 95 | |
| Study design | | | | | |
| Parallel | 10 | 10.74 (8.97–12.51) | <0.001 | 37 | 0.370 |
| Crossover | 16 | 12.01 (8.32–15.70) | <0.001 | 93 | |
| Hormone | | | | | |
| Dual | 3 | 22.00 (10.30–33.69) | < 0.001 | 92 | 0.048 |
| Single | 23 | 10.12 (7.83–12.42) | <0.001 | 87 | |
| Timing of intervention* | 10 | | | | 0 700 |
| 24 h | 19 | 11.14 (8.49–13.78) | <0.001 | 89 | 0.793 |
| Overnight Church duration mounths | 18 | 14.79 (10.96–18.62) | <0.001 | 87 | |
| Study duration, months | 14 | 12.01 (6.71.10.00) | <0.001 | 03 | 0.225 |
| <1 | 14 | 12.81 (6.71–18.90) | <0.001 | 93 | 0.235 |
| 21 Maan and warm | 12 | 9.95 (8.63–11.28) | <0.001 | 49 | |
| | 10 | | <0.001 | 00 | 0.655 |
| <14 | 18 | 11.45 (8.55-14.57) | < 0.001 | 90 | 0.055 |
| ≥14 Comparator | 0 | 11.29 (0.94–15.03) | <0.001 | 85 | |
| Comparator | 0 | 10 22 /7 20 12 42) | <0.001 | 80 | 0 5 6 0 |
| | 0 | 10.32 (7.20 - 13.43) | < 0.001 | 05 | 0.569 |
| Mixed | 15 | 13.36 (7.11-19.05) | < 0.001 | 91 | |
| Mixed | 5 | 12.01 (9.27-14.70) | <0.001 | 42 | |
| TBR (<3.9 mmol/L) | | | | | |
| All comparisons | 24 | -0.59 (-1.02 to -0.15) | 0.008 | 86 | |
| Setting | | | | | |
| Unsupervised | 17 | -0.23 (-0.67 to 0.21) | 0.308 | 87 | 0.004 |
| Supervised | 7 | -2.03 (-3.09 to -0.96) | <0.001 | 52 | |
| Study design | | | | | |
| Parallel | 10 | -0.45 (-0.86 to -0.03) | 0.034 | 77 | 0.400 |
| Crossover | 14 | -0.87 (-1.64 to -0.11) | 0.025 | 85 | |
| Hormone | | | | | |
| Dual | 3 | -1.98 (-2.94 to -1.02) | <0.001 | 0 | 0.063 |
| Single | 21 | -0.42 (-0.86 to 0.03) | 0.064 | 86 | |
| Timing of intervention* | 10 | | 0.000 | 70 | 0.012 |
| 24 n | 18 | -0.63 (-1.05 to -0.21) | 0.003 | 79 | 0.913 |
| Overnight Study dynation months | 13 | -0.65 (-1.46 to 0.17) | 0.121 | 86 | |
| | 12 | 1.04 (1.97 + 0.20) | 0.015 | 69 | 0.196 |
| <1 >1 | 15 | -1.04 (-1.87 to -0.20) | 0.015 | 00 | 0.180 |
| | 11 | -0.28 (-0.81 to 0.24) | 0.294 | 91 | |
| | 17 | $0.46(0.93 \pm 0.01)$ | 0.054 | 84 | 0 769 |
| <14 >1 <i>1</i> | 17 | -0.46(-0.55(0)0.01) -0.79(-1.61 to 0.03) | 0.054 | 04 75 | 0.769 |
| Comparator | 1 | -0.75 (-1.01 to 0.05) | 0.000 | 75 | |
| SAP | 13 | -0.52 (-1.13 to 0.10) | 0 098 | 88 | 0 227 |
| CSIL or MDI | 8 | -1.12(-2.02 to -0.22) | 0.014 | 58 | 0.227 |
| Mixed | 3 | -0.59 (-1.02 to -0.15) | 0.334 | 30 | |
| | | 0.00 (1.02 to 0.12) | 0.001 | | |
| TAR (>10 mmol/L) | | | | | |
| All comparisons | 19 | -12.19 (-14.65 to -9.73) | <0.001 | 83 | |
| Setting | | | | | |
| Unsupervised | 14 | -9.74 (-11.60 to -7.88) | < 0.001 | 64 | 0.005 |
| Supervised | 5 | -19.32 (-26.64 to -12.00) | <0.001 | 85 | |
| Study design | _ | | | | |
| Parallel | 7 | -9.41 (-11.88 to -6.95) | < 0.001 | 47 | 0.153 |
| Crossover | 12 | -14.04 (-17.78 to -10.30) | <0.001 | 88 | |
| Hormone | 2 | 10.21 / 20.24 / 0.20 | -0.001 | 67 | 0.424 |
| Duai Cin ele | 3 | -18.31 (-28.24 to -8.38) | <0.001 | 87 | 0.121 |
| Single | 16 | -11.21 (-13.68 to -8.73) | <0.001 | 81 | |
| iming of intervention* | 14 | 12 02 / 14 76 to 0 20) | <0.001 | 05 | 0.000 |
| | 14 | -12.03 (-14.76 t0 -9.30) | < 0.001 | 85 | 0.098 |
| Overnight | 10 | -17.54 (-22.90 to -11.78) | 0.002 | 63 | |

Continued on p. 2304

| SubgroupComparisons, nRD (95% Cl)P l^2 , % $P_{interac}$ Study duration, months </th <th>Table 1—Continued</th> <th></th> <th></th> <th></th> <th></th> <th></th> | Table 1—Continued | | | | | |
|---|------------------------|----------------|---------------------------|---------|--------------------|--------------|
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | Subgroup | Comparisons, n | RD (95% CI) | Р | l ² , % | Pinteraction |
| $\begin{array}{cccccccccccccccccccccccccccccccccccc$ | Study duration, months | | | | | |
| $ \begin{array}{c ccccccccccccccccccccccccccccccccccc$ | <1 | 9 | -17.14 (-21.89 to -12.38) | < 0.001 | 76 | 0.002 |
| Mean age, years <14 13 $-12.46 (-15.50 \text{ to } -9.43)$ <0.001 86 0.76 ≥ 14 6 $-11.64 (-16.29 \text{ to } -6.99)$ <0.001 75Comparator | ≥1 | 10 | -8.67 (-10.24 to -7.11) | < 0.001 | 48 | |
| $\begin{array}{c ccccccccccccccccccccccccccccccccccc$ | Mean age, years | | | | | |
| ≥14 6 -11.64 (-16.29 to -6.99) <0.001 75 | <14 | 13 | -12.46 (-15.50 to -9.43) | < 0.001 | 86 | 0.765 |
| Comparator | ≥14 | 6 | -11.64 (-16.29 to -6.99) | < 0.001 | 75 | |
| comparator | Comparator | | | | | |
| SAP 11 -11.97 (-15.21 to -8.73) <0.001 85 0.70 | SAP | 11 | -11.97 (-15.21 to -8.73) | < 0.001 | 85 | 0.704 |
| CSII or MDI 6 -13.22 (-19.51 to -6.93) <0.001 85 | CSII or MDI | 6 | -13.22 (-19.51 to -6.93) | < 0.001 | 85 | |
| Mixed 2 -11.88 (-18.63 to -5.13) 0.001 73 | Mixed | 2 | -11.88 (-18.63 to -5.13) | 0.001 | 73 | |

RD, risk difference. *If the study reported both 24-h and overnight results, both results were extracted.

except one, which was gray literature from ClinicalTrials.gov. In 22 trials, a singlehormone AID system was assessed, mostly compared with SAP. Two trials assessed a dual-hormone AID system by comparing it with CSII. The remaining study was a threeway crossover trial, which assessed dualhormone AID, single-hormone AID, and CSII.

In two studies assessing AID compared with SAP. the control treatment comprised SAP combined with low glucose suspend. Among trials evaluating singlehormone AID systems, four used the Diabetes Assistant platform, four used the Florence configuration, four used the Control-IQ system, two used the MiniMed 670G, two used Florence M or CamAPS FX configuration, and the remainder used other systems, mixed systems, or did not provide relevant details. In terms of setting, six trials were conducted in a diabetes camp, one was conducted in a hotel, and 18 were conducted in participants' home. At baseline, participant mean age ranged from 3.9 to 17 years, and HbA_{1c} ranged from 7.3 to 10.6%. Characteristics of individual studies are summarized in Supplementary Table 1. Most examined items were assessed as low or unclear risk except blinding. Because of the nature of the intervention, blinding for participants and personnel seemed to be impracticable. The detailed risk-of-bias assessment is available in Supplementary Table 2. The quality of evidence for each outcome was rated following the GRADE framework (Supplementary Table 3).

Primary Outcome

All meta-analysis results are presented as pooled effect estimates for AID systems versus conventional insulin therapy. Twenty-six comparisons from 25 studies with 1,345 participants were pooled for the primary outcome of TIR. Compared with conventional insulin therapy, use of AID was associated with an increased percentage of time (164 additional minutes) in target range (MD 11.38% [95% CI 9.01–13.76], P < 0.001; high certainty) (Fig. 1). There was high statistical heterogeneity (l^2 = 89%). Effect sizes of individual studies ranged from an MD of -6.30 to 33.00%.

The primary outcomes stratified by type of AID system, timing of intervention, study design, mean age, study duration, and comparator are shown in Table 1. Of note, use of AID had a significant favorable effect on TIR in all subgroups. The favorable effect was consistent when AID was used over 3 months (nine trials, MD 10.45% [95% CI 8.71-12.20]) or 6 months (five trials, MD 10.87% [7.11-14.63]). AID systems seem to have a greater improvement in TIR in supervised settings (seven trials, MD 15.25% [5.55-24.96]) compared with unsupervised settings (18 trials, MD 9.88% [8.05-11.72]), but the test for interaction was not significant (Pinteraction = 0.156).

We repeated the meta-analysis for the primary outcome using only one comparison from the study that used a three-way crossover design, and the result was almost unchanged. Additionally, we performed a post hoc sensitivity analysis by excluding studies that enrolled participants aged >18 years, and the primary outcome was also consistent (20 trials, MD 11.56% [95% CI 8.85-14.28]). There was no significant publication bias among the studies based on Egger test that compared AID with conventional insulin therapy (P = 0.201), and the funnel plot did not show evidence of publication bias visually (Supplementary Fig. 2).

Secondary Outcomes

Twenty-four comparisons with 1,287 participants were pooled for TBR (<3.9 mmol/L). TBR (<3.9 mmol/L) was 0.59% (95% CI -0.59 to -0.15, P = 0.008; low certainty) lower for AID systems compared with conventional insulin therapy (Fig. 2). AID systems showed improvement in TBR (<3.9 mmol/L) in supervised settings (six trials, MD -2.03%[-3.09 to -0.96], P < 0.001) but not in unsupervised settings (17 trials, MD -0.23% [-0.67 to 0.21], P = 0.308), and the subgroup difference was significant (P_{interaction} = 0.004). The favorable effect of AID use for TBR (<3.9 mmol/L) over conventional insulin therapy was significant in dual-hormone, short-term, and 24-h subgroups, but not in single-hormone, long-term, and overnight subgroups. Twenty-one comparisons with 1,051 participants were pooled for TBR (<3.0 mmol/L). There was no difference in TBR (<3.0 mmol/L) between AID and conventional insulin therapy (MD -0.09% $[-0.19 \text{ to } 0.01], P = 0.088, I^2 = 82\%; \text{ low}$ certainty).

Nineteen comparisons with 1,032 participants were pooled for TAR (>10 mmol/L). TAR (>10 mmol/L) was 12.19% (95% CI −14.65 to −9.73, P < 0.001; high certainty) shorter by AID use compared with conventional insulin therapy. Use of AID had a significant favorable effect on TAR (>10 mmol/L) in all subgroups. AID systems had a greater improvement in TAR (>10 mmol/L) in supervised settings (four trials, MD -19.32% [-26.64 to -12.00], P < 0.001) compared with unsupervised settings (14 trials, MD -9.74% [-11.60 to -7.88], P < 0.001) (P_{interaction} = 0.005). Long-term study duration was associated with lower improvement in TAR (>10 mmol/L) compared with short-term

| Study | Patients | Timing of intervention | Study duration | | Mean difference with 95% Cl | Weight (%) |
|--|------------------------|--------------------------|----------------|-----------|--------------------------------|---------------|
| Unsupervised | | | | | | |
| Kariyawasam 2022 | 21 | 24 h | 6 weeks | | -2.62 [-4.78, -0.46] | 2.55 |
| Abraham 2021 | 135 | 24 h | 6 months | - | -1.90 [-2.50, -1.30] | 6.12 |
| Sharifi 2016 | 12 | Overnight | 4 days | | -1.80 [-3.43, -0.17] | 3.49 |
| Isganaitis 2021 | 63 | 24 h | 6 months | | -0.70 [-1.10, -0.30] | 6.55 |
| Breton 2020 | 101 | 24 h | 16 weeks | | -0.40 [-0.81, 0.01] | 6.54 |
| Russell 2018 | 20 | 24 h | 5 days | | -0.40 [-1.65, 0.85] | 4.39 |
| Messer 2022 | 165 | 24 h | 13 weeks | | -0.30 [-0.60, -0.00] | 6.72 |
| Wadwa 2023 | 102 | 24 h | 13 weeks | - | -0.20 [-0.75, 0.35] | 6.24 |
| Forlenza 2018 | 28 | Overnight | 21 days | | 0.00 [-0.99, 0.99] | 5.09 |
| Ware 2022a | 74 | 24 h | 16 weeks | | 0.10 [-0.35, 0.55] | 6.45 |
| McVean 2023 | 113 | 24 h | 12 months | | 0.20 [-0.30, 0.70] | 6.35 |
| Hovorka 2014 | 16 | Overnight | 3 weeks | | 0.20 [-2.70, 3.10] | 1.68 |
| Tauschmann 2016b | 12 | 24 h | 3 weeks | | 0.40 [-1.07, 1.87] | 3.86 |
| Ware 2022b | 133 | 24 h | 6 months | | 0.53 [-1.77, 2.83] | 2.34 |
| Thabit 2015 | 25 | Overnight | 12 weeks | | 0.83 [0.59, 1.07] | 6.80 |
| Tauschmann 2016a | 12 | 24 h | 7 days | _ | 1.20 [-0.37, 2.77] | 3.63 |
| Boughton 2022 | 97 | 24 h | 24 months | | 2.80 [-0.60, 6.20] | 1.31 |
| Heterogeneity: $\tau^2 = 0.5$ | 5, $I^2 = 86.839$ | %, H ² = 7.59 | | • | -0.23 [-0.67, 0.21] | |
| Test of $\theta_i = \theta_j$: Q(16) = | 121.51, p = 0 | 0.00 | | | | |
| Supervised | | | | | | |
| Del Favero 2016 | 30 | 24 h | 3 days — | | -4.70 [-6.59, -2.81] | 2.98 |
| Russell 2016 | 19 | 24 h | 5 days | | -3.20 [-5.61, -0.79] | 2.20 |
| Russell 2014 | 32 | 24 h | 5 days | | -1.80 [-3.33, -0.27] | 3.72 |
| Haidar 2015b | 33 | Overnight | 3 days | | -1.70 [-3.14, -0.26] | 3.93 |
| Breton 2017 | 32 | 24 h | 5 days | | -1.40 [-3.09, 0.29] | 3.38 |
| DeBoer 2017 | 12 | 24 h | 3 days | | -0.50 [-2.76, 1.76] | 2.40 |
| Haidar 2015a | 33 | Overnight | 3 days | | 0.00 [-3.49, 3.49] | 1.26 |
| Heterogeneity: $\tau^2 = 1.0$ | 2, $I^2 = 51.69^\circ$ | %, H ² = 2.07 | | • | -2.02 [-3.09, -0.96] | |
| Test of $\theta_i = \theta_j$: Q(6) = 1 | 2.42, p = 0.0 | 5 | | | | |
| Overall | | | | • | -0.58 [-1.02, -0.15] | |
| Heterogeneity: $\tau^2 = 0.7$ | $0, I^2 = 85.989$ | %, H ² = 7.13 | | | | |
| Test of $\theta_i = \theta_j$: Q(23) = | 164.02, p = 0 | 0.00 | | | | |
| | | | _ | | | |

Figure 2—Forest plot for TBR (<3.9 mmol/L) by study included in the meta-analysis. Full reference citations are available in the Supplementary Material.

($P_{\text{interaction}} = 0.002$). Twelve comparisons with 824 participants were pooled for TAR (>13.9 mmol/L), and AID showed favorable effects compared with conventional insulin therapy (MD -4.14% [-6.29 to -1.99], P < 0.001, $l^2 = 92\%$; low certainty).

Most trials reported no serious adverse effects. Episodes of severe hypoglycemia were mentioned in six studies (23–28), and there was no distinct difference between AID systems and conventional insulin therapy. Only four diabetic ketoacidosis events were reported in three trials (23,28,29), all of which occurred in the AID groups.

DISCUSSION

This meta-analysis included 25 randomized crossover trials involving 1,345 patients and provided an overview of glycemic control of AID systems compared with conventional insulin therapy in outpatient settings in children and adolescents with T1D. The use of AID systems resulted in an 11.38% (95% CI 9.01–13.76) increased TIR, equivalent to 164 min/day. This finding was also verified by its effect on TAR (176 min less than control treatment) and TBR (8 min less). The primary outcome was robust and consistent in all sensitivity analyses performed. This favorable effect was observed consistently when the AID was used continuously over 3 and 6 months, which was not analyzed in previous studies. Most trials reported no serious adverse events, such as severe hypoglycemia and diabetic ketoacidosis.

This study also suggested greater improvement in TIR for dual-hormone compared with single-hormone AID systems. The characteristics of different AID systems and conventional insulin treatments have been reviewed comprehensively in three reviews (9,30,31). A dual-hormone system has been shown to be superior to a single-hormone system in improving nocturnal glucose control in children and adolescents with T1D in one three-way crossover trial (32). However, dual-hormone systems have only been tested for 3-5 days with very close supervision (diabetes camp) in a small number of children. Future, freerange studies are required in which dualhormone and single-hormone AID systems are compared directly with each other in children and adolescents.

AID systems reduced TBR only by 0.59% compared with conventional insulin therapy (P = 0.008); however, significant heterogeneity was present across trials. The subgroup analysis showed that the reduction was significant in dual-hormone systems and supervised settings but not in single-hormone systems and unsupervised settings. This finding contrasts with previous meta-analyses of AID use that did not have age limitations (10,12) but is consistent with a recent meta-analysis focused on young people (14). More trials of dual-hormone AID in young people are needed.

Despite heterogeneity in interventions and comparators used, our systematic review provides a valid and up-to-date overview about the use of AID in children and adolescents. Previous metaanalyses of AID in all age-groups showed favorable effects both overnight and over a 24-h period, but the longest follow-up was 12 weeks (10,12). The favorable effect was also evident in a subgroup analysis for the pediatric population (12), which was consistent with our meta-analysis. However, our study included more trials, performed comprehensive assessment, and had greater reliability. Furthermore, Karageorgiou et al. (13) analyzed the effectiveness of AID in the nonadult population (aged <18 years), and the results suggested

that AID systems are superior to standard SAP treatment for T1D. However, only 5 of 19 included studies were in outpatient settings, and 4 trials were not randomized. The effect of AID systems in children and adolescents was examined in a recent systematic review and meta-analysis of 26 RCTs (915 participants) (14). However, one-half of the trials included had a follow-up duration of <5 days, and the validity and clinical interpretation were undermined by methodological decisions regarding the analysis of crossover trials, date extraction of overnight glucose control, and inadequate subgroup analysis (14).

Initially, we used a correlation coefficient of 0.5 to calculate the mean and SE of the paired differences if there was no reporting in crossover trials. A post hoc validation for the correlation coefficient was done using a method recommended by the Cochrane Collaboration (33), which yielded a value of 0.8 for TIR in the largest crossover trial (27). Therefore, the result of our meta-analysis was more conservative.

Our study has several strengths. The meta-analysis followed the PRISMA guidelines and a protocol that was registered with PROSPERO. We conducted a comprehensive search of multiple databases and included all available RCTs of AID compared with conventional insulin therapy. Risk of bias for included trials was assessed using a valid methodological tool, and quality of evidence for each outcome was evaluated using GRADE. Subgroup analyses were performed to explain heterogeneity, and sensitivity analyses were conducted to examine the robustness of the results.

We also acknowledge some limitations. First, the sample size was small in most trials, which reduced the precision of effect estimates. Second, most included trials were considered at high risk of performance bias because of infeasibility of blinding patients and physicians to the allocation assignments. Third, statistical assumptions were made in this study. We retrieved means and SDs from medians and IQRs, respectively, which would be most problematic within the secondary analysis of TBR (<3.0 mmol/L) and TAR (>13.9 mmol/L). Additionally, a correlation coefficient of 0.5 was assumed to calculate the mean and SE of the paired differences. Fourth, heterogeneity was high in most analyses, which could be attributed to differences in study design, duration of intervention, continuous

glucose monitoring systems, AID algorithms, insulin pumps, and persisting impact of human factors. Finally, the results of this meta-analysis might not apply to some clinically relevant subgroups, such as those with increased hypoglycemia burden, hypoglycemia unawareness, and high HbA_{1c}. While some studies have investigated the association between AID use and HbA_{1c} (34,35) and hypoglycemia unawareness (36), further studies are warranted to fully clarify these relationships.

In conclusion, this systematic review and meta-analysis shows that AID systems are more effective than conventional insulin therapy for children and adolescents with T1D in outpatient settings. AID systems increase TIR both in short-term and long-term intervention.

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